

1. Upper bound on $\hat{Q}_a(t)$ for many samples

Suppose ν is a bandit model with 1-sub-gaussian arms. Show that under the UCB Algorithm $\hat{Q}_a(t) < Q_a + \Delta_a$ with probability $1 - \delta$, given that $T_a(t) > \frac{2\log(1/\delta)}{\Delta_a^2}$.

Solution:

Proof: Consider w.l.o.g. that $\mathbb{P}^\pi(T_a(t) = n) > 0$ for all $n \in \{1, \dots, t - (k - 1)\}$. (UCB chooses each of the $k - 1$ suboptimal arms at least once in the beginning). First we can observe that $T_a(t) > \frac{2\log(1/\delta)}{\Delta_a^2}$ is equivalent to $\Delta_a > \sqrt{\frac{2\log(1/\delta)}{T_a(t)}}$, so we will now consider the probability of $\hat{Q}_a(t) - Q_a \geq \sqrt{\frac{2\log(1/\delta)}{T_a(t)}}$. Then, considering the intersection with the condition $T_a(t) = n$ for some $n \leq t - (k - 1)$ yields

$$\begin{aligned} & \mathbb{P}^\pi\left(\hat{Q}_a(t) - Q_a \geq \sqrt{\frac{2\log(1/\delta)}{T_a(t)}} \cap (T_a(t) = n)\right) \\ &= \mathbb{P}^\pi\left(\frac{1}{T_a(t)} \sum_{i=1}^t X_i \mathbf{1}_{\{A_i=a\}} - Q_a \geq \sqrt{\frac{2\log(1/\delta)}{T_a(t)}} \cap (T_a(t) = n)\right) \\ &= \mathbb{P}^\pi\left(\frac{1}{n} \sum_{i=1}^t X_i \mathbf{1}_{\{A_i=a\}} - Q_a \geq \sqrt{\frac{2\log(1/\delta)}{n}} \cap (T_a(t) = n)\right) \\ &= \mathbb{P}^\pi\left(\frac{1}{n} \sum_{i=1}^t X_i \mathbf{1}_{\{A_i=a\}} - Q_a \geq \sqrt{\frac{2\log(1/\delta)}{n}} \mid T_a(t) = n\right) \mathbb{P}^\pi(T_a(t) = n) \\ &\leq \delta \mathbb{P}^\pi(T_a(t) = n). \end{aligned}$$

Note that a conditional probability is still a probability measure so we can use the normal Hoeffdings inequality in the last step.

Furthermore, we obtain that

$$\begin{aligned}
& \mathbb{P}^\pi \left((\hat{Q}_a(t) < Q_a + \Delta_a) \mid (T_a(t) > \frac{2 \log(1/\delta)}{\Delta_a^2}) \right) \\
& \geq \mathbb{P}^\pi \left(\hat{Q}_a(t) - Q_a < \sqrt{\frac{2 \log(1/\delta)}{T_a(t)}} \mid (T_a(t) > \frac{2 \log(1/\delta)}{\Delta_a^2}) \right) \\
& = \mathbb{P}^\pi \left(\hat{Q}_a(t) - Q_a < \sqrt{\frac{2 \log(1/\delta)}{T_a(t)}} \mid \bigcup_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} (T_a(t) = n) \right) \\
& \geq \frac{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi \left(\hat{Q}_a(t) - Q_a < \sqrt{\frac{2 \log(1/\delta)}{n}} \cap (T_a(t) = n) \right)}{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n)} \\
& = \frac{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n) - \mathbb{P}^\pi \left(\hat{Q}_a(t) - Q_a \geq \sqrt{\frac{2 \log(1/\delta)}{n}} \cap (T_a(t) = n) \right)}{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n)} \\
& \geq \frac{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n) - \delta \mathbb{P}^\pi (T_a(t) = n)}{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n)} \\
& = 1 - \delta,
\end{aligned}$$

where we used the definition of conditional expectation and that $\mathbb{P}(A \cap B) = \mathbb{P}(B) - \mathbb{P}(A^c \cap B)$.

2. Regret bounds for UCB on σ -subgaussian bandit models

For σ -subgaussian bandit models the UCB exploration bonus is modified as

$$\text{UCB}_a(t) := \begin{cases} \infty, & T_a(t) = 0, \\ \hat{Q}_a(t) + \sqrt{\frac{2\sigma^2 \log(\frac{1}{\delta})}{T_a(t)}}, & T_a(t) \neq 0. \end{cases}$$

Check that the regret bound in Theorem 1.3.8 using $\delta = \frac{1}{n^2}$ changes to

$$R_n(\pi) \leq 3 \sum_{a \in \mathcal{A}} \Delta_a + 16\sigma^2 \log(n) \sum_{a: Q_a \neq Q_*} \frac{1}{\Delta_a},$$

and that this leads to

$$R_n(\pi) \leq 8\sigma \sqrt{Kn \log(n)} + 3 \sum_{a \in \mathcal{A}} \Delta_a$$

in Theorem 1.3.9.

Solution:

The general idea of the proof of 1.3.8 does not change. We start by defining G_m in the same way except that now

$$G_{2,m} := \left\{ \omega : \bar{Q}_m^{(a)}(\omega) + \sqrt{\frac{2\sigma^2 \log(\frac{1}{\delta})}{m}} < Q_{a_1} \right\}.$$

The proof that $G_m \subseteq H_m$ works exactly the same using the modified notion of G_m . Similarly, $\mathbb{P}(G_1^c) \leq n\delta$ still holds. We then choose $m = \left\lceil \frac{2\sigma^2 \log(1/\delta)}{1/4\Delta_a^2} \right\rceil$ in order to assure

$$\Delta_a - \sqrt{\frac{2\sigma^2 \log(\frac{1}{\delta})}{m}} \geq \frac{1}{2}\Delta_a.$$

By Hoeffdings inequality for σ -subgaussian random variables we obtain similarly to 1.3.8 that $\mathbb{P}(G_{2,m}^c) \leq \exp(-\frac{m\Delta_a^2}{8})$ and finally the regret bound in the exercise. As in 1.3.9, rewriting

$$R_n(\pi) \leq n\Delta + 3 \sum_{a \in \mathcal{A}} \Delta_a + \frac{16K\sigma^2 \log(n)}{\Delta}$$

and optimizing over Δ leads to the alternative regret bound stated in the exercise.

3. Best Baseline

The variance of a random vector X is defined by to be $\mathbb{V}[X] := \mathbb{E}[\|X\|_2^2] - \|E[X]\|_2^2$. Show by differentiation that

$$b_* = \frac{\mathbb{E}_{\pi_\theta}[X_A \|\nabla \log \pi_\theta(A)\|_2^2]}{\mathbb{E}_{\pi_\theta}[\|\nabla \log \pi_\theta(A)\|_2^2]}$$

is the baseline that minimises the variance of the unbiased estimators

$$(X_A - b)\nabla \log(\pi_\theta(A)), \quad A \sim \pi_\theta,$$

of $\nabla J(\theta)$.

Solution:

We have

$$\begin{aligned} & \mathbb{V}\left((X_A - b)\nabla \log(\pi_\theta(A))\right) \\ &= \mathbb{E}\left[(X_A - b)^2 \|\nabla \log(\pi_\theta(A))\|_2^2\right] - \left\| \mathbb{E}\left[(X_A - b)\nabla \log(\pi_\theta(A))\right] \right\|_2^2 \\ &= \mathbb{E}\left[(X_A - b)^2 \|\nabla \log(\pi_\theta(A))\|_2^2\right] - \left\| \mathbb{E}\left[X_A \nabla \log(\pi_\theta(A))\right] \right\|_2^2, \end{aligned}$$

where we used the baseline trick in the last equation. We define $f(A) = \|\nabla \log(\pi_\theta(A))\|_2$ to have a better overview. Then

$$\begin{aligned} & \mathbb{V}\left((X_A - b)\nabla \log(\pi_\theta(A))\right) \\ &= \mathbb{E}\left[(X_A - b)^2 f(A)^2\right] - \left\| \mathbb{E}\left[X_A \nabla \log(\pi_\theta(A))\right] \right\|_2^2 \\ &= \mathbb{E}\left[X_A^2 f(A)^2\right] - 2b\mathbb{E}\left[X_A f(A)^2\right] + b^2\mathbb{E}\left[f(A)^2\right] - \left\| \mathbb{E}\left[X_A \nabla \log(\pi_\theta(A))\right] \right\|_2^2. \end{aligned}$$

We calculate the first derivative as

$$\begin{aligned} & \frac{\partial \mathbb{V}\left((X_A - b)\nabla \log(\pi_\theta(A))\right)}{\partial b} \\ &= -2\mathbb{E}\left[X_A f(A)^2\right] + 2b\mathbb{E}\left[f(A)^2\right]. \end{aligned}$$

Solving for the root gives

$$b^* = \frac{\mathbb{E}\left[X_A f(A)^2\right]}{\mathbb{E}\left[f(A)^2\right]},$$

which is a minimum, as the second derivative $2\mathbb{E}\left[f(A)^2\right] \geq 0$ almost surely. Plugging in the definition of f proves the claim.