

3. Excercise Sheet

1. Upper bound on $\hat{Q}_a(t)$ for many samples

Suppose ν is a bandit model with 1-sub-gaussian arms. Show that under the UCB Algorithm $\hat{Q}_a(t) < Q_a + \Delta_a$ with probability $1 - \delta$, given that $T_a(t) > \frac{2 \log(1/\delta)}{\Delta_a^2}$.

Solution:

Proof: Consider w.l.o.g. that $\mathbb{P}^\pi(T_a(t) = n) > 0$ for all $n \in \{1, \dots, t - (k - 1)\}$. (We just go up to $t - (k - 1)$ because we have to choose $k - 1$ times a different arm as every arm has to be played once in the beginning.) First we obtain that $T_a(t) > \frac{2 \log(1/\delta)}{\Delta_a^2}$ is equivalent to $\Delta_a > \sqrt{\frac{2 \log(1/\delta)}{T_a(t)}}$.

So we will now first consider the probability of $\hat{Q}_a(t) - Q_a \geq \sqrt{\frac{2 \log(1/\delta)}{T_a(t)}}$. Then, we first consider the intersection with condition $T_a(t) = n$ for some $n \leq t - (k - 1)$.

$$\begin{aligned} & \mathbb{P}^\pi \left(\hat{Q}_a(t) - Q_a \geq \sqrt{\frac{2 \log(1/\delta)}{T_a(t)}} \cap (T_a(t) = n) \right) \\ &= \mathbb{P}^\pi \left(\frac{1}{T_a(t)} \sum_{i=1}^t X_i \mathbf{1}_{\{A_i=a\}} - Q_a \geq \sqrt{\frac{2 \log(1/\delta)}{T_a(t)}} \cap (T_a(t) = n) \right) \\ &= \mathbb{P}^\pi \left(\frac{1}{n} \sum_{i=1}^t X_i \mathbf{1}_{\{A_i=a\}} - Q_a \geq \sqrt{\frac{2 \log(1/\delta)}{n}} \cap (T_a(t) = n) \right) \\ &= \mathbb{P}^\pi \left(\frac{1}{n} \sum_{i=1}^t X_i \mathbf{1}_{\{A_i=a\}} - Q_a \geq \sqrt{\frac{2 \log(1/\delta)}{n}} \mid T_a(t) = n \right) \mathbb{P}^\pi(T_a(t) = n) \\ &\leq \delta \mathbb{P}^\pi(T_a(t) = n). \end{aligned}$$

Note that a conditional probability is still a probability measure so we can use the normal Hoeffdings inequality in the last step.

Further we obtain that

$$\begin{aligned}
& \mathbb{P}^\pi \left((\hat{Q}_a(t) < Q_a + \Delta_a) \mid (T_a(t) > \frac{2 \log(1/\delta)}{\Delta_a^2}) \right) \\
& \geq \mathbb{P}^\pi \left(\hat{Q}_a(t) - Q_a < \sqrt{\frac{2 \log(1/\delta)}{T_a(t)}} \mid (T_a(t) > \frac{2 \log(1/\delta)}{\Delta_a^2}) \right) \\
& = \mathbb{P}^\pi \left(\hat{Q}_a(t) - Q_a < \sqrt{\frac{2 \log(1/\delta)}{T_a(t)}} \mid \bigcup_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} (T_a(t) = n) \right) \\
& \geq \frac{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi \left(\hat{Q}_a(t) - Q_a < \sqrt{\frac{2 \log(1/\delta)}{n}} \cap (T_a(t) = n) \right)}{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n)} \\
& = \frac{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n) - \mathbb{P}^\pi \left(\hat{Q}_a(t) - Q_a \geq \sqrt{\frac{2 \log(1/\delta)}{n}} \cap (T_a(t) = n) \right)}{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n)} \\
& \geq \frac{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n) - \delta \mathbb{P}^\pi (T_a(t) = n)}{\sum_{n=\lceil \frac{2 \log(1/\delta)}{\Delta_a^2} \rceil}^{t-(k-1)} \mathbb{P}^\pi (T_a(t) = n)} \\
& = 1 - \delta,
\end{aligned}$$

where we used the definition of conditional expectation and that $\mathbb{P}(A \cap B) = \mathbb{P}(B) - \mathbb{P}(A^c \cap B)$.

2. Best Baseline

The variance of a random vector X is defined by to be $\mathbb{V}[X] := \mathbb{E}[\|X\|_2^2] - \|E[X]\|_2^2$. Show by differentiation that

$$b_* = \frac{\mathbb{E}_{\pi_\theta} [X_A \|\nabla \log \pi_\theta(A)\|_2^2]}{\mathbb{E}_{\pi_\theta} [\|\nabla \log \pi_\theta(A)\|_2^2]}$$

is the baseline that minimises the variance of the unbiased estimators

$$(X_A - b) \nabla \log(\pi_\theta(A)), \quad A \sim \pi_\theta,$$

of $\nabla J(\theta)$.

Solution:

We have

$$\begin{aligned}
& \mathbb{V}\left((X_A - b)\nabla \log(\pi_\theta(A))\right) \\
&= \mathbb{E}\left[(X_A - b)^2 \|\nabla \log(\pi_\theta(A))\|_2^2\right] - \left\|\mathbb{E}\left[(X_A - b)\nabla \log(\pi_\theta(A))\right]\right\|_2^2 \\
&= \mathbb{E}\left[(X_A - b)^2 \|\nabla \log(\pi_\theta(A))\|_2^2\right] - \left\|\mathbb{E}\left[X_A \nabla \log(\pi_\theta(A))\right]\right\|_2^2,
\end{aligned}$$

where we used the baseline trick in the last equation. We define $f(A) = \|\nabla \log(\pi_\theta(A))\|_2$ to have a better overview. Then

$$\begin{aligned}
& \mathbb{V}\left((X_A - b)\nabla \log(\pi_\theta(A))\right) \\
&= \mathbb{E}\left[(X_A - b)^2 f(A)^2\right] - \left\|\mathbb{E}\left[X_A \nabla \log(\pi_\theta(A))\right]\right\|_2^2 \\
&= \mathbb{E}\left[X_A^2 f(A)^2\right] - 2b\mathbb{E}\left[X_A f(A)^2\right] + b^2\mathbb{E}\left[f(A)^2\right] - \left\|\mathbb{E}\left[X_A \nabla \log(\pi_\theta(A))\right]\right\|_2^2
\end{aligned}$$

We calculate the first derivative

$$\begin{aligned}
& \frac{\partial \mathbb{V}\left((X_A - b)\nabla \log(\pi_\theta(A))\right)}{\partial b} \\
&= -2\mathbb{E}\left[X_A f(A)^2\right] + 2b\mathbb{E}\left[f(A)^2\right].
\end{aligned}$$

Solving for the root gives

$$b^* = \frac{\mathbb{E}\left[X_A f(A)^2\right]}{\mathbb{E}\left[f(A)^2\right]},$$

which is a minimum, as the second derivative $2\mathbb{E}\left[f(A)^2\right] \geq 0$ almost surely. Plugging in the definition of f proves the claim.